ORIGINAL ARTICLE

Multifaceted allergen avoidance during infancy reduces asthma during childhood with the effect persisting until age 18 years

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ABSTRACT

Background Asthma is a chronic disease that often starts in childhood. The key risk factors are a child’s environment and their genetic characteristics. The aim of this study was to evaluate the efficacy of environmental modification in the first 12 months of life on the prevalence of asthma in high-risk individuals.

Methods Children (n=120) considered at high risk of allergic disorders (either dual heredity or single heredity and a high cord total IgE), were enrolled in a single-blinded, randomised controlled trial. Infants in the intervention arm were either breast fed with the mother on a low allergen diet or given an extensively hydrolysed formula. Exposure to house dust mite allergen was reduced. The control group followed standard advice. Children were assessed at ages 1, 2, 4, 8 and 18 years for the presence of asthma and atopy.

Results At 18 years of age, there was a significantly lower prevalence of asthma in the prevention group compared with the control group (OR: 0.23, 95% CI 0.08 to 0.70, p=0.01), primarily due to asthma that developed during childhood but persisted until age 18 years. Repeated-measure analysis showed that there was an overall reduction in asthma prevalence from 1 to 18 years (OR: 0.51, CI 0.32 to 0.81, p=0.04). Prevalence of atopy was not significantly different between the two groups at age 18.

Conclusion Comprehensive allergen avoidance in the first year of life is effective in preventing asthma onset in individuals considered at high risk due to heredity. The effect occurs in the early years, but persists through to adulthood.

INTRODUCTION

Over the last 50 years, the prevalence of asthma has increased dramatically, with an estimated global occurrence of 300 million.1 Despite intensive efforts to develop novel therapeutic agents, asthma is still an incurable disease with pharmacotherapy at best achieving abeyance of symptoms. Research has highlighted the importance of the interaction of gene and environment, particularly in the early years of life.2 Atopy is arguably the most significant single risk factor for asthma with a population-attributable risk of 56% in some, but not all populations.3 Environmental factors represent an opportunity for intervention in asthma prevention.4 House dust mite exposure (HDM),5 dietary intake,6 microbial7 and viral exposures8 are potential factors influencing asthma onset. Optimal timing of primary prevention is unknown, but early life immunological development in atopic infants,9 and evidence of airway changes consistent with asthma in infants,10 argues for early intervention within the first few months of life. Further, early sensitisation to HDM has been consistently associated with childhood asthma.11

We hypothesised that in infants genetically predisposed to atopy, allergen avoidance of both house dust mites and common food allergens during infancy may lead to a reduction in the development of allergic diseases with the benefit continuing beyond the actual period of avoidance. Consequently, 120 infants were recruited into this intervention study in 1990 and assessed at the ages of 1, 2, 4 and 8 years. Outcomes at these follow-ups have been reported previously and have shown significant reduction in asthma, eczema and atopy in the intervention group.12–16 We now report the outcome of this randomised controlled trial examining the preventive effect of a multifaceted allergen avoidance strategy in the first year of life.

METHODS

Detailed descriptions of the intervention methodology have been published previously.12 Briefly, in 1990, 120 infants who were considered at high risk of developing allergic disease on the basis of dual heredity (two or more immediate family members (parents and/or sibling) with an allergic disorder) or...
single heredity (one parent or sibling with an allergic disorder) plus cord total IgE >0.5 kU/l were recruited at birth. Participants were randomised into a prevention group (n=58) and a control group (n=62) using computer-generated random allocation numbers.12 The study was single-blinded, so that the participants’ grouping was concealed from the researchers until after assessment at each follow-up. From birth until 12 months, lactating mothers and infants in the prevention arm followed a diet of strict avoidance of: dairy products, egg, soya, fish and shellfish, peanut and tree nuts. HDM reduction measures were taken, including vinyl mattress covers and use of acaricide in bedrooms and living rooms. Dust samples collected at 0 and 9 months showed that the prevention group had significantly lower levels of HDM than the control group (median Der p1 5.86 µg/g (interquartile range 3.55–10.93) vs 15.31 µg/g (7.25–27.49), p<0.0001). The control group was provided with standard recommendations prevalent at the time. A 100% follow-up was achieved at 1, 2, 4 and 8 years of age. Assessments consisted of parental questionnaires regarding symptoms of asthma and allergic diseases. Skin prick testing was performed in the majority of participants at each follow-up.12–15

### Eighteen-year follow-up
The study was approved by Southampton & South West Hampshire Research Ethics Committee (07/HO504/188) and the protocol registered with ISRCTN (96472018). Of the original 120 participants, 114 (95%) gave informed consent for the assessment at age 18. The researchers assessing the participants were blinded to their group allocation. As in previous assessment, all participants answered the International Study of Asthma and Allergies in Childhood (ISAAC) core questionnaire17 and, where appropriate, reported the type of asthma medication used and answered the Juniper asthma-specific quality-of-life questionnaire.

For 103 participants (48 in the prevention group and 55 in the control group), allergy skin prick testing was performed by a standardised method16 to house dust mite (*Dermatophagoides pteronyssinus*).
Acholine bronchial challenge was performed in accordance with baseline prior to spirometry. Baseline pulmonary function was measured using Koko spirometry software (PDS Instrumentation, Longmont, USA) according to standardised methodology. Methacholine bronchial challenge was performed in accordance with international guidelines. In order to compare bronchial hyperresponsiveness where participants did not achieve a 20% (or more) drop in their FEV1, the dose response slope was calculated. Induction, processing and analysis of sputum samples were undertaken in accordance with ERS guidelines.

**Definitions**

A participant was defined as atopic where (s)he had at least one positive skin prick test to a food or aeroallergen. A skin prick test was defined as positive where the mean weal diameter was \( \geq 3 \) mm larger than the negative control. Asthma was defined as a positive response to (1) has a doctor diagnosed you with asthma and, either (2) have you wheezed in the last 12 months or (3) are you on inhaled corticosteroids.

**Table 1 Asthma, asthma phenotypes and atopy at 18-year follow-up**

<table>
<thead>
<tr>
<th></th>
<th>Prevention (n = 56)*</th>
<th>Control (n = 58)*</th>
<th>p Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>Any current asthma</td>
<td>6 (10.7)</td>
<td>15 (25.9)</td>
<td>0.04</td>
</tr>
<tr>
<td>Persistent asthma</td>
<td>3 (5.4)</td>
<td>10 (17.2)</td>
<td>0.04</td>
</tr>
<tr>
<td>Late-onset asthma</td>
<td>3 (5.4)</td>
<td>5 (8.6)</td>
<td>0.38</td>
</tr>
<tr>
<td>Remitted asthma</td>
<td>7 (12.5)</td>
<td>5 (8.6)</td>
<td>0.50</td>
</tr>
<tr>
<td>Ever had asthma</td>
<td>13 (23.2)</td>
<td>20 (34.5)</td>
<td>0.18</td>
</tr>
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</table>

For any current asthma, n (%) = 56 for both prevention and control; for persistent asthma, n (%) = 58 for both prevention and control. For late-onset asthma, n (%) = 56 for prevention and n (%) = 58 for control. For remitted asthma, n (%) = 56 for prevention and n (%) = 58 for control. For ever had asthma, n (%) = 56 for prevention and n (%) = 58 for control.

Asthma in childhood is variable in terms of onset, remission and relapse. To test whether the intervention was associated with the different types, we classified participants into one of the following groups: persistent asthma (asthma onset at/ before 8 years and current asthma at 18 years), late-onset asthma (asthma at 18 years but no prior history of asthma), remitted asthma (no asthma at 18 but prior history of asthma) or never asthma (no current or prior history of asthma).

**Statistical methods**

Details of statistical methodology are provided in the on-line supplement. Briefly, an intention-to-treat analysis was performed using all available data. The sample size was limited to the original 120 participants who were recruited prenatally 18 years ago. The primary outcome of the current analysis was asthma at 18 years, and longitudinally from 1 to 18 years of age.

Depending on the significance in bivariate analysis, six explanatory variables were tested for inclusion in the binary regression model: (1) group, (2) dual heredity, (3) family history of asthma (at least one parent or sibling with asthma), (4) whether or not the subject was a firstborn child, (5) exposure to smoke in the 2 years preceding the 18-year follow-up and (6) maternal smoking during pregnancy.

To assess the relationship between groups and the different types of asthma (never asthma, persistent asthma, remitted asthma and late-onset asthma), a multinomial logistic regression model was built using the same model-building strategy as for the binary regression model. Longitudinal analysis was undertaken using generalised estimating equations (GEE) with a logit link function and an independent correlation structure. The GEE analysis was based on the 114 subjects who were followed-up at 18 years, as well as at age 1, 2, 4 and 8 (a total of 547 data points).

**RESULTS**

**Asthma at 18 years: cross-sectional analysis**

At the age of 18 years, 114 of 120 (95%) were assessed; 56/58 (96.6%) from the prevention group and 58/62 (93.5%) from the control group (figure 1). The prevalence of asthma was significantly lower in the prevention group compared with the control group (10.7% and 25.9%, respectively); the OR was 0.34, 95% CI 0.12 to 0.96, \( p=0.04 \) (table 1). The significantly lower prevalence of asthma at 18 years was due mainly to a lower prevalence of persistent asthma rather than late-onset asthma.

The binary logistic regression found that only two variables were significantly associated with asthma at 18 years once other explanatory variables were held constant: group and family history of asthma. The final model contained these two covariates. The odds of asthma at age 18 were 4.33 times greater if there was a family history of asthma (CI 1.37 to 13.74, \( p=0.01 \)). Once family history of asthma was held constant, the odds of asthma at age 18 for the prevention group were 0.23 times the odds for the control group (CI 0.08 to 0.70, \( p=0.01 \)). In other words, the effect of the intervention was stronger once we adjusted for family history of asthma. We tested for interaction between group and family history of asthma, but this was not significant (\( p=0.48 \)).

The frequency and severity of asthma symptoms and asthma-specific quality-of-life scores were not significantly different between the groups (table 2). Further, there was no significant difference in terms of lung function, bronchial hyperresponsiveness, FeNO or Airways inflammatory cells.

**Table 3 shows the results of the multinomial logistic regression model.** The figures in the table are predicted probabilities,

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**Asthma**

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that is, the proportion of the study participants in each group who would have each type of asthma according to the model (thus, each row sums to 100%). These results confirm that the association between the intervention and asthma at 18 is almost entirely due to the prevention group being significantly less likely to have persistent asthma; there was no significant association between group and late-onset or remitted asthma, even when family history of asthma was held constant. The predicted probability of persistent asthma among those with no family history of asthma was five times higher in the control group than in the prevention group (8.5% and 1.6%, respectively). Similarly, the predicted probability of persistent asthma among those with a family history of asthma was four times higher in the control group than in the prevention group (8.5% and 1.6%, respectively). The difference between the groups was noticeable from the earliest follow-up (OR: 0.42, CI 0.25 to 0.83, p = 0.007 (figure 3)). Males was significantly associated with an increased risk of asthma within this model (OR 1.71, CI 1.11 to 2.6, p = 0.02), and no other variable was significantly associated with an increased risk of asthma within this model.

Repeated-measure analysis showed that over the length of follow-up there was a significantly lower period of prevalence atopy in the prevention group compared with the control group (OR: 0.42, CI 0.25 to 0.83, p = 0.007) (figure 3). The difference between the groups was noticeable from the earliest follow-up at 1 year, and persisted through to 8 years, when the difference had narrowed and became non-significant at 18 years (figure 3).

**DISCUSSION**

Our study demonstrates a significant and sustained reduction in the prevalence of asthma in participants who underwent a comprehensive food and house dust mite allergen-avoidance strategy in the first year of life. Atopy is a significant factor in the development of asthma. Our analysis demonstrates a significant reduction in the prevalence of atopy and, specifically, HDM sensitisation in the prevention group in early childhood (table 1), resulting in a significant reduction in persistent asthma, and therefore, the difference in asthma prevalence still being significant at 18 years, specifically atopic asthma. There was no significant effect on atopy or asthma developing during adolescence. However, the longitudinal analysis demonstrates an overall significant difference in the prevalence of both asthma and atopy over the duration of follow-up (figures 2 and 3). Further follow-up will determine whether our intervention has successfully prevented the onset of asthma or merely delayed it to later adult life.

The study has several limitations. The number of randomised participants was small, and so the study may not have had adequate power for all the variables tested. However, the finding of a significant difference in the primary outcome measure of asthma prevalence is a valid one, as power is only critical to prevent Type II error where the null hypothesis, of no significant difference between the two groups, is inappropriately accepted. Another limitation is that participants were aware of the group allocation and, hence, reporting bias remains a possibility. However, preventive effect has been consistent throughout 18 years, including asthma treatment prescribed by their physicians. In view of the open design, the control group may have taken measures to reduce allergen level, but that would serve to reduce the difference between the groups. The active intervention stopped at 12 months; however, it is possible that parents in the

**Table 2** Symptoms, quality of life, lung function and airway inflammation in participants with asthma

<table>
<thead>
<tr>
<th>Variable</th>
<th>Control group, no family history of asthma</th>
<th>Control group, family history of asthma</th>
<th>Prevention group, no family history of asthma</th>
<th>Prevention group, family history of asthma</th>
<th>p Value</th>
</tr>
</thead>
<tbody>
<tr>
<td>FEV1/FVC, % predicted</td>
<td>94.71 (16.23)</td>
<td>101.75 (7.95)</td>
<td>0.20</td>
<td></td>
<td></td>
</tr>
<tr>
<td>FVC, % predicted</td>
<td>91.80 (12.43)</td>
<td>96.50 (17.03)</td>
<td>0.55</td>
<td></td>
<td></td>
</tr>
<tr>
<td>FEV1, % predicted</td>
<td>95.15 (21.77)</td>
<td>100.88 (17.24)</td>
<td>1.00</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Sputum neutrophils, %</td>
<td>9.5 (2.9)</td>
<td>5.3 (2.9)</td>
<td>0.08</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Sleep affected &gt;1 night a week</td>
<td>1 (16.7%)</td>
<td>2 (12.5%)</td>
<td>0.91</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Wheeze on exertion</td>
<td>6 (100.0%)</td>
<td>13 (86.7%)</td>
<td>1.00</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Wheeze affecting speech</td>
<td>2 (33.3%)</td>
<td>5 (33.3%)</td>
<td>1.00</td>
<td></td>
<td></td>
</tr>
<tr>
<td>&gt;4 wheeze attacks</td>
<td>4 (66.7%)</td>
<td>5 (33.3%)</td>
<td>0.36</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Asthma quality-of-life scores</td>
<td>5.06 (4.00–6.10)</td>
<td>6.01 (5.60–6.50)</td>
<td>0.07</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

**Table 3** Results of multinomial logistic regression model: predicted probability of being in each ‘asthma’ group

<table>
<thead>
<tr>
<th>Predicted probability of:</th>
<th>Never asthma</th>
<th>Remitted asthma</th>
<th>Persistent asthma</th>
<th>Late-onset asthma</th>
</tr>
</thead>
<tbody>
<tr>
<td>Prevention group, no family history of asthma</td>
<td>90.3%</td>
<td>6.1%</td>
<td>1.6%*</td>
<td>2.1%</td>
</tr>
<tr>
<td>Control group, no family history of asthma</td>
<td>80.2%</td>
<td>6.1%</td>
<td>8.5%*</td>
<td>5.2%</td>
</tr>
<tr>
<td>Prevention group, family history of asthma</td>
<td>70.9%</td>
<td>15.3%</td>
<td>7.0%*</td>
<td>6.8%</td>
</tr>
<tr>
<td>Control group, family history of asthma</td>
<td>47.5%</td>
<td>11.7%</td>
<td>28.0%*</td>
<td>12.9%</td>
</tr>
</tbody>
</table>

*p < 0.05.
prevention group may have continued with some allergen-avoidance measures. There was concern regarding an under-reporting of asthma symptoms by adolescents in the intervention group if they were aware of their group allocation. We asked 23 randomly selected participants regarding their group allocation. Only three could correctly identify their group. Thus, we do not think that response bias influenced the outcome of this study.

The two groups were matched in terms of education, smoking and participation in this study (table E1). While there were significantly more participants with dual heredity in the prevention group and firstborns in the control group, neither of these variables influenced the outcome in either cross-sectional or longitudinal analysis. We did not adjust for multiple testing as the primary outcome variables, that is asthma and atopy, were determined a priori, before this analysis, at the inception of this RCT, and subsequently at every follow-up.12–16 Symptoms, lung function and markers of airway inflammation did not significantly differ between patients with asthma in the prevention and control groups (table 2). However, the numbers were small and the possibility of type II error (false negative) cannot be excluded. All we can say is that the intervention reduced the prevalence of asthma globally rather than preventing onset at any single age.

We found a significant reduction in asthma using the dual intervention of HDM avoidance and dietary modification. A larger, multicentre trial with larger sample sizes and genetic homogeneity, the prevention effect seen in this trial. For a number of reasons, which include small sample sizes and genetic homogeneity, the prevention effect seen in this study may not be generalisable. However, given the heterogeneous nature of asthma, it is unlikely that a single intervention of any kind would be effective in all participants. The significance of asthma prevention is such that further studies are warranted to identify the subgroups where reduction in allergen exposure might be effective.

In conclusion, our study provides evidence that a combined dietary and environmental allergen-avoidance strategy in the first year of life is successful as a primary prevention strategy for asthma in high-risk individuals, with benefits persisting into early adulthood. There is an urgent need to replicate these findings in a large multicentre study with stratification to investigate effectiveness and cost-effectiveness, and the immunological mechanisms underlying this approach.

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**Figure 2** Prevalence of asthma over 18 years of follow-up. Generalised estimating equation-repeated measures analysis, adjusted for first born status, dual heredity, exposure to paternal and maternal smoking and pets in the home.

**Figure 3** Prevalence of atopy over 18 years of follow-up. Generalised estimating equation-repeated measures, bivariate analysis.
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Contributors  SHA conceived the idea and worked with all the authors on the plan for recruitment, assessments and analysis, and reviewed and revised draft. MS undertook the recruitment and assessments, performed data entry and initial statistical analysis and wrote the first draft, AN reviewed and revised the statistical analysis, GR, SM and RK helped with the draft revisions. All the authors reviewed and discussed the results, contributed to the manuscript and approved the final version.

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