Comparing and combining studies of bronchial responsiveness

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Background: There is no standardised protocol for the measurement of bronchial responsiveness. Results from different studies are difficult to compare and combine.

Methods: Analyses are divided between those of a continuous outcome, which can be directly standardised as effect size, and those based on a binary outcome. A published method is used to convert an odds ratio to equivalent effect size.

Results: The use of effect size allows comparison between studies using a continuous outcome but different protocols, provided the relevant standard deviation is reported. Effect size from a continuous outcome and that derived from an odds ratio from an equivalent analysis gave similar results.

Conclusions: Systematic reviews which include both continuous effect estimates and odds ratios can include both in one meta-analysis, provided relevant standard deviations are published for the former. Authors are encouraged to report these in all fields in which measurement protocols vary.

The measurement of bronchial responsiveness (BHR) is advocated in population studies of asthma. While not synonymous with asthma, it is an objective measure that has particular advantages in multicultural studies. BHR has also been used in the diagnosis of asthma and the measurement of severity in clinical settings.

The results of a scientific study are of little use unless they are in a form that is useful to the readers and allows comparison with previous or future work. This normally requires that the measurement protocols are similar. While some protocol variation exists for many measurements, this is a particular problem for BHR.

Results of histamine or methacholine challenge are most commonly summarised in clinical studies by the dose (PD$_{20}$) or concentration (PC$_{20}$) that produces a 20% fall in forced expiratory volume in one second (FEV$_1$). A logarithmic scale is considered appropriate for the analysis, and results are frequently expressed in doubling doses or concentrations. For example, in a randomised controlled trial of two treatments for asthma, the result for BHR might be one doubling dose difference—that is, that after one treatment, compared to the other, twice as much of the provoking agent was required to produce a 20% fall in FEV$_1$, or, equivalently, that BHR was reduced by one doubling dose.

Consequences of Protocol Variation for Clinical Studies

Results reported as PD$_{20}$, PD$_{10}$, PC$_{20}$, etc can be converted to doubling dose units, and authors of meta-analyses have used this method to combine results from different studies. This does not necessarily overcome the problem of protocol variation. The response to treatment may be more variable in one study than in another as a result of the protocol differences. Provided a standard deviation of the outcome on the doubling dose or other logarithmic scale is reported, this can be used to calculate a standardised effect for each study. The estimate and its associated standard error are divided by the standard deviation to give “effect size”. This does not affect the associated p value. When the estimate is a difference in means, effect size is also known as standardised difference. This approach was used by Abramson et al to combine results from 12 studies on immunotherapy in asthma with non-specific BHR as outcome and 14 studies that measured allergen specific BHR.

A Further Problem in Population Studies

In clinical studies on asthmatic patients, each participant is likely to have a measurable PD$_{20}$ on each occasion. In a population study the maximum dose of provocation agent will be limited by safety considerations, and at least 50% of the population will have a less than 20% fall in FEV$_1$, at the highest dose administered. For these participants, PD$_{20}$ is said to be “censored”. There are three ways of dealing with this in the analysis.

One is to use a method known as censored regression, implemented in at least one major statistical program. This produces estimates on the log(PD$_{20}$) scale, but is heavily dependent on the assumption that log(PD$_{20}$) is normally distributed. However, there is evidence for this assumption, and it has been used in the analysis of several population studies of the relation of BHR to urinary electrolytes.

A second method is to use an alternative summary statistic for BHR. Several measures of “slope” have been proposed which have the advantage that a value can be defined for each person. The simplest of these, due to O’Connor, is the fall in FEV$_1$, divided by the final dose. Provided a standard deviation is reported, estimates using censored regression of PD$_{20}$ or a slope outcome can be combined or compared with each other using effect size as described above.

A third method is the most common in population studies. Participants are simply divided into those “with” and those “without” BHR according to whether PD$_{20}$ is less than the maximum dose. Descriptive statistics are proportions of the sample “with BHR”, analysis is using logistic regression, and results are given as odds ratios. This analysis makes few assumptions and is relatively simple to understand. However, it is wasteful of information. In addition, BHR in the population has a unimodal distribution, so any division into subjects “with” or “without” BHR is purely arbitrary and usually determined by the maximum dose administered.

Cook and Strachan found 10 studies that reported an odds ratio for BHR for children exposed to environmental tobacco smoke compared with those who were not. Odds ratios can be combined on the log scale and antilogged to give a final...
In the relation of BHR to sodium excretion, giving a regression coefficient of log10(PD20) = 1.750, which is the average of two separate meta-analyses. This allows for the fact that the logistic coefficient is the natural logarithm of the odds ratio associated with a unit increase in log10 (sodium) per 10-fold increase; a more useful quantity might be the change in BHR associated with a 50% increase which can be derived by multiplying each estimate by log10 (1.5).

Table 1 shows effect sizes derived from results in a paper that reported BHR summarised as a slope measure, devised for use in the European Community Health Survey, and analysed by multiple regression (table 6 in reference), and equivalent results from logistic regression (table 7 in reference). The residual standard deviation from the multiple regression was 2.10 and the variation explained 33.3%, so the residual standard deviation was calculated as √(0.667/2.10). As stated, the sign of each effect size derived from the multiple regression was reversed so that it was compatible with the comparable logistic regression coefficient.

Table 2 shows effect sizes derived from results in a paper that reported BHR summarised as a slope measure, devised for use in the European Community Health Survey, and analysed by multiple regression (table 6 in reference), and equivalent results from logistic regression (table 7 in reference).

<table>
<thead>
<tr>
<th>Factor</th>
<th>Effect size derived from multiple regression*</th>
<th>Effect size derived from odds ratio from logistic regression</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Effect size 95% confidence interval p value</td>
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</tr>
<tr>
<td>Total IgE (log kU/l)</td>
<td>0.08 -0.05 to 0.20 0.23</td>
<td>0.16 -0.06 to 0.38 0.16</td>
</tr>
<tr>
<td>Specific IgE to:</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Cat</td>
<td>0.34 0.05 to 0.63 0.02</td>
<td>0.55 0.11 to 1.00 0.02</td>
</tr>
<tr>
<td>House dust mite (HDM) titre [just positive]</td>
<td>0.42 0.22 to 0.62 &lt;0.001</td>
<td>0.43 0.10 to 0.75 0.01</td>
</tr>
<tr>
<td>per 10 kU/l increase in HDM titre above 0.35 kU/l</td>
<td>0.13 0.03 to 0.23 0.01</td>
<td>0.09 0.06 to 0.24 0.22</td>
</tr>
<tr>
<td>Timothy grass</td>
<td>0.29 0.08 to 0.49 0.005</td>
<td>0.64 0.33 to 0.95 &lt;0.001</td>
</tr>
<tr>
<td>Cladosporium</td>
<td>0.84 0.37 to 1.30 &lt;0.001</td>
<td>1.20 0.49 to 1.91 0.001</td>
</tr>
<tr>
<td>Birch</td>
<td>-0.28 -0.62 to 0.05 0.09</td>
<td>-0.74 -1.33 to 0.17 0.01</td>
</tr>
<tr>
<td>FEV1 as standardised difference from predicted value</td>
<td>-0.22 -0.31 to -0.12 &lt;0.001</td>
<td>-0.35 -0.52 to -0.18 &lt;0.001</td>
</tr>
<tr>
<td>FEV1 %FVC</td>
<td>-0.05 -0.07 to -0.04 &lt;0.001</td>
<td>-0.06 -0.08 to -0.03 &lt;0.001</td>
</tr>
</tbody>
</table>

Table 1 Example of conversion of results in the relation of BHR to sodium excretion in two different forms to effect size

Table 2 Effect sizes for the relation of bronchial hyperresponsiveness (BHR) to serum total IgE, individual allergens, and lung function derived from a multiple regression analysis of a continuous measure compared with those expressed as odds ratios from the equivalent logistic regression model

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For more information, please refer to the original articles and references cited in the text.
greater p values due to the loss of information on dichotomising BHR. Where the two effect sizes do appear to differ, which was notable for the association with specific IgE to Timothy grass and, to a lesser extent, with specific IgE to birch, this is due not to a problem with the conversion of odds ratio to effect size but to the fact that the multiple regression and logistic regression did in this case give different results, with the p value for the latter less than that for the former.

**DISCUSSION**

The above method for conversion of a ln(odds ratio) to effect size shows that the two are essentially equivalent. However, an analysis of a continuous outcome is generally more powerful than one based on an arbitrary division of the scale into two groups. This is not an argument for a slope measure of BHR over PD20, as the existence of a value for each subject does not automatically imply more information. Although the result implies that an odds ratio may be little affected by the alteration of cut off point, other statistics will change. Peat et al reported that BHR has high specificity for asthma and low sensitivity, but a greater maximum dose of provoking agent and cut off point would increase sensitivity and decrease specificity.

The method was illustrated using results from Burney et al and Chinn et al because, in each study, the data were analysed in the two ways; in the former the residual standard deviation from the linear regression was reported and in the latter it could be calculated from published results, enabling the calculation of effect size in both cases. More often results are reported without the standard deviation or only as an odds ratio. When a mixture of odds ratios and estimates based on continuous outcomes are to be compared or combined, it is essential to obtain an estimate of the residual standard deviation for the latter. Publication of residual standard deviations from every analysis of variance and multiple linear regression should therefore be encouraged. This also applies to any outcome in addition to BHR for which a completely standardised protocol does not exist.

A result expressed as effect size is not, of course, as useful clinically as one in doubling dose units. An alternative would be to assume an underlying standard deviation for studies reporting just an odds ratio and use this to convert the effect size to approximate doubling dose units. Again, only if authors report standard deviations will it be possible to know what standard deviation might be assumed.

It seems unlikely that a standardised protocol can be achieved for BHR, so there will be a continuing need to compare and combine results in the above manner. While any comparison may require this, clearly the need is greatest in conducting a systematic review and meta-analysis. Conducting two meta-analyses, one of odds ratios and the other of standard deviations from every analysis of variance and multiple linear regression would therefore be encouraged. This automaticaly implies that the two are essentially equivalent. However, an analysis of a continuous outcome is generally more powerful than one based on an arbitrary division of the scale into two groups. This is not an argument for a slope measure of BHR over PD20, as the existence of a value for each subject does not automatically imply more information. Although the result implies that an odds ratio may be little affected by the alteration of cut off point, other statistics will change. Peat et al reported that BHR has high specificity for asthma and low sensitivity, but a greater maximum dose of provoking agent and cut off point would increase sensitivity and decrease specificity.

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