Usual dietary salt intake and asthma in children: a case-control study

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Abstract

Background – A decline in host resistance due to an alteration in diet – primarily of salt – was recently put forward as a possible explanation for rising rates of asthma. Methods – A case-control study was conducted in participants in a prevalence survey which included 187 children with asthma (defined by prior diagnosis and/or a decline in forced expiratory volume in one second (FEV₁) of ≥10% after exercise) and 145 age and sex matched controls. Subjects were selected from 989 children aged 5–13 years attending 18 elementary schools on the island of Montreal. Usual dietary salt intake was estimated from a food frequency questionnaire administered to the mother, and a salt intake score was used to group the children into quartiles from I (lowest) to IV (highest salt intake). Bronchial hyperresponsiveness to methacholine was assessed by Yan’s method. Cases and controls were combined in one group to examine the relationship of salt intake to bronchial hyperresponsiveness to methacholine. Methacholine responsiveness was expressed as a dose-response slope and ranks of dose–response slopes were used in the analysis. Results – After accounting for important confounding variables, there was no association between asthma and salt intake, while methacholine dose-response slope ranks increased with increasing salt intake and methacholine responsiveness was greater in the highest quartile than in the lowest quartile of salt intake. The median dose-response slopes in % fall in FEV₁ per μmol methacholine for quartiles I, II, III, and IV were 5.4, 5.9, 7.7, and 8.7. Conclusions – No association was found between asthma or exercise-induced bronchospasm and dietary salt intake. Bronchial hyperresponsiveness to methacholine did, however, appear to increase with greater salt intake, but the relevance of this association to asthma is unclear.

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Keywords: salt intake, bronchial hyperresponsiveness, children.

Asthma morbidity and mortality is greater in communities adopting a more western lifestyle and in migrants as they move from rural underdeveloped to urban westernised areas. The lack of an adequate epidemiological explanation for this phenomenon, coupled with the increase in salt consumption with urbanisation, led Burney to hypothesise that increased sodium intake might be partly responsible for the increased mortality and morbidity due to asthma. He tested the hypothesis using regional data from England and Wales and found a strong correlation between table salt purchases and asthma mortality in men and children of both sexes, but not in women. Two controlled crossover trials have recently found increasing dietary salt intake to result in worsening asthma symptoms and an increase in bronchodilator consumption. The degree to which such salt loading might be applicable to salt consumption observed in uncontrolled situations is unclear.

In contrast, Lieberman and Heimer in a non-blinded randomised crossover trial involving 17 asthmatic patients found no significant difference in peak flows or peak flow variability between periods of low and high salt intake. Pistelli et al. conducted a community based cross sectional study in children aged 9–16 years and found the reported use of table salt to be strongly related to respiratory symptoms suggestive of asthma in boys, but not in girls. These discrepancies in the reported results of the different investigations and the limited data arising from community based studies of children led us to examine the relationship of normal dietary salt intake to asthma in children.

Methods

A case-control study among participants in a cross sectional survey was conducted in the city of Montreal, Canada from April 1990 to November 1992. Eighteen schools were selected on the island of Montreal in order to represent a broad range of socioeconomic status. One class from each school from each of grades 1 (age 5–7 years), 3 (age 8 and 9 years), and 5 (age 10–13 years) was selected. A questionnaire on history of asthma diagnosed by a doctor, respiratory symptoms and illnesses, asthma in a parent, exposure to second-hand tobacco smoke, housing conditions including the current presence of pets, and a letter of consent were completed by parents.

In the school gymnasium each subject’s age, sex, height, and weight were recorded and spirometric tests were carried out sitting and with nose clips using two Collins 10 litre water-
sealed spirometers (Warren E Collins, Brain-tree, Massachusetts, USA) according to current American Thoracic Society guidelines, and the best FEV, from any flow-volume curve, both at baseline and after exercise, was used for analysis.

Following five minutes of inactivity, heart rate was measured using a digital plethysmograph (Heart Rate Inc, Costa Mesa, California, USA). Children then ran around the gymnasium for six minutes at a pace judged sufficient to attain \( \geq 90\% \) of the predicted maximal heart rate; heart rate was remeasured immediately on completion of exercise. Five and 10 minutes after completion of exercise the spirometric tests were repeated on the same spirometer. Average room temperature and relative humidity, as well as percentage maximum heart rate attained, were similar among children who did and did not demonstrate a \( \geq 10\% \) fall in FEV, after exercise (exercise-induced bronchospasm). Subjects did not take part in the exercise test if they were excused from gym class or if their FEV, was below 70\% predicted. No specific instructions were given concerning the use of medications including those for asthma. The exercise test was completed successfully in 989 children. Only 11 (1-11\%) had been dispensed a bronchodilator (for example, \( \beta\) agonist) at the time of the test.

The subjects for the case-control study were selected from the participants in the cross sectional survey. A case was defined by the presence of either (1) a history of asthma as reported by parents, or (2) a fall in FEV, of \( \geq 10\% \) after exercise. The next child on the alphabetical list of the same sex as the case, but without either of these criteria for asthma, was chosen as a control. If the parents of the case refused further participation, both this child and the one chosen to serve as a control were not studied further, while if a child chosen as a control (that is, without asthma or exercise-induced bronchospasm) refused, the next appropriate child on the list was selected. For 38 cases a comparison child could not be obtained.

A food frequency questionnaire designed specifically to assess the normal intake of salt in the diet of the child was administered to a parent (almost always the mother). Food items high in sodium (4) that are frequently consumed were grouped into nine food groupings; portion sizes were not indicated. Parents were also asked about their child's preference for salty foods. Parents chose from four frequency categories for each of the nine food groups their child might consume, and a score was obtained from the sum of the items multiplied by the frequency of each item (maximum value 27).

The salt preference score obtained by asking the question "Does your child like salty foods?" (very much, 2; somewhat, 1; not at all, 0) was then added to obtain the overall salt intake score (maximum value 29). The questionnaire is provided in the appendix. Food frequency is widely used as the dietary method of choice to provide an approximate ranking of the normal intake of individuals on a specific dietary component. In measuring salt intake, such an estimate of usual intake is particularly important because of the high day-to-day variability in intake.

The last occupation of the parents was transformed into the corresponding codes of the Canadian Classification and Dictionary of Occupations. These codes were then converted into socioeconomic status scores for the child based on income and education level for each occupation from the tables developed by Blisken and colleagues. The highest score from either parent was retained for analysis.

In order to investigate in detail the risk factors for bronchial hyperresponsiveness, methacholine bronchoprovocation and allergy skin tests were performed on cases and controls at home in the evening. All children with an FEV, of \( >75\% \) of their forced vital capacity had a methacholine challenge test using the method of Yan. A long protocol with nine incremental doses was followed among the cases, with six incremental doses given to the controls. The test began with three inhalations of 0-9\% saline with FEV, being measured one minute later. Provided that the FEV, did not fall by more than 10\% of the baseline value, methacholine solutions were administered according to the following cumulative doses in \( \mu\)mol: 0-030, 0-060, 0-124, 0-244, 0-499, 0-996, 1-990, 3-913, and 7-80 for the long protocol and 0-030, 0-094, 0-477, 1-967, 3-89, and 7-78 for the short protocol. At each dose level subjects performed inspiratory capacity inhalation with a five second breath hold while seated and wearing nose clips. A forced expiratory manoeuvre was performed 60 seconds after each dose. The challenge was stopped when the FEV, had fallen by 20\% or more, or the final dose had been reached.

Solutions for allergy skin testing comprised histamine (1 mg/ml), normal saline, Dermatophagoides pteronyssinus, Dermatophagoides farinae, mixed grass pollens, tree pollens, ragweed, mixed moulds, Aspergillus sp, cat epithelium, and cockroach antigen. Needles (26 gauge) were used to break the skin. The resultant weal diameters were measured at right angles using Vernier callipers at 10 and 15 minutes. A weal of \( \geq 3\) mm was taken as a positive response. If there was no positive response to histamine, skin testing was regarded as invalid.

The study was approved by the ethics committee of McGill University.

**DATA ANALYSIS**

The salt intake score was used to group the children into quartiles (quartile I with the lowest and quartile IV with the highest salt score, respectively). The relationship between asthma and salt intake was examined using a logistic regression model before and after adjusting for potential confounding variables. Statistical significance of the regression coefficients was determined by the \( \chi^2 \) approximation to the likelihood ratio statistic.

Cases and controls were combined into one group in order to investigate the relationship of salt intake to bronchial hyperresponsiveness to methacholine. For the methacholine broncho-
provocation test the dose-response data of
each child were summarised according to the
method of O’Connor et al21 - that is, percentage
decline in FEV₁/dose - where percentage de-
cline in FEV₁ was defined as the decline in
FEV₁ (from the post saline value) after the final
methacholine dose, and dose was defined as the
final cumulative methacholine dose ad-
imistered. This method of expressing metha-
choline responsiveness in population studies
has been validated by several groups.22,23
The dose-response slope was ranked because of
the non-normal distribution which persisted,
despite various transformations. The ranks
were used as an outcome variable for sub-
sequent multivariate linear regression analysis
where approximate normality was observed in
the residuals.
Statistical analysis was carried out using SAS
and EGRET statistical softwares.24,25

Results
A total of 1274 children were eligible to par-
ticipate in the cross sectional survey from 18
Montreal schools selected. Of these, the parents
of 130 children (10-2%) refused participation
for their children, while a further 75 children
(5-9%) did not return the questionnaire and
consent form. There were no meaningful
differences between participants and non-par-
ticipants as to the age of the child (mean (SD)
8-8 (1-8) versus 8-0 (1-9)), sex (boys: 51% versus
55%), race (Caucasians: 78% versus 80-9%), socioeconomic status assessed by neigh-
bourhood census data (poorest socioeconomic
status quartile: 27% versus 23%). Spirometric
tests were not performed for a further 23 chil-
dren because of illness or absence from school.
Of the 1046 children who attempted spirom-
metric testing, 28 (2-7%) were unable to com-
plete the test successfully and such failure was
more common among younger children but
was unrelated to reported respiratory illness,
symptoms, or socioeconomic status. The spi-
rometric data of a further 28 children (2-7%) were lost after the test. One child (0-1%) was
excluded because of a severe attack of asthma
at the time of the test.
Of the 989 children (77-6%) who par-
ticipated in the cross sectional survey and had
spirometric data, 269 children (27-2%) met
the case definition and for 231 of these cases
a child of the same sex in the same class was
available to serve as a control. Of the 269

children with either exercise-induced broncho-
spasm or a history of asthma (80 of whom had
a history of asthma alone), 187 (70%) were
visited at home. Of the 231 children without
exercise-induced bronchospasm or a history of
asthma, 145 (63%) were visited. Among the
cases (table 1), non-participants were more
likely to be Caucasian, have mothers or fathers
who currently smoke, or mothers who had
smoked during pregnancy. Asthma, but not a
history of other respiratory illness, was more
common among participants, indicating an in-
creased interest in the study by parents of
children with this common childhood disorder
despite the fact that our study was presented
as an inquiry into respiratory health in general.
Among the control group differences between
participants and non-participants were less pro-
nounced, although they were mostly in a similar
direction to the cases.

The prevalences of wheeze, allergy skin test
positivity, and bronchial hyperresponsiveness
to methacholine were higher among the cases
than the controls, while pre-exercise lung func-
tion levels were lower (table 2).

Table 1  Descriptive characteristics of the subsample selected for further study at home

<table>
<thead>
<tr>
<th></th>
<th>Cases</th>
<th>Controls</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Participants</td>
<td>Non-participants</td>
</tr>
<tr>
<td></td>
<td>(n = 187)</td>
<td>(n = 82)</td>
</tr>
<tr>
<td>Boys (%)</td>
<td>55-4</td>
<td>53-2</td>
</tr>
<tr>
<td>Caucasians (%)</td>
<td>74-6</td>
<td>86-1</td>
</tr>
<tr>
<td>Mother smokes (%)</td>
<td>37-3</td>
<td>45-2</td>
</tr>
<tr>
<td>Father smokes (%)</td>
<td>43-5</td>
<td>50-8</td>
</tr>
<tr>
<td>Smoked while pregnant (%)</td>
<td>29-6</td>
<td>45-1</td>
</tr>
<tr>
<td>History of asthma (%)</td>
<td>44-9</td>
<td>37-2</td>
</tr>
<tr>
<td>Other respiratory history* (%)</td>
<td>47-1</td>
<td>58-9</td>
</tr>
</tbody>
</table>

* History of either pneumonia, bronchiolitis, whooping cough, or cough.
NA = not applicable.

Table 2  Indicators of asthma among the cases and controls

<table>
<thead>
<tr>
<th></th>
<th>Cases (n = 187)</th>
<th>Controls (n = 145)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Ever wheeze (%)</td>
<td>49-2</td>
<td>19-7</td>
</tr>
<tr>
<td>Allergy skin test positivity (%)**</td>
<td>43-7</td>
<td>31-4</td>
</tr>
<tr>
<td>Asthma in a parent (%)</td>
<td>18-5</td>
<td>6-0</td>
</tr>
<tr>
<td>Methacholine responsive (%)†</td>
<td>31-0</td>
<td>19-4</td>
</tr>
<tr>
<td>PEV:</td>
<td>-1-50</td>
<td>+1-54</td>
</tr>
<tr>
<td>FEV₁/FVC</td>
<td>-0-40</td>
<td>+0-39</td>
</tr>
</tbody>
</table>

FEV₁ = forced expiratory volume in one second; FVC = forced vital capacity.
** > 3 mm weal to one or more of the nine aeroallergens tested (see text for explanation).
† Children with >20% fall in FEV₁, after <7-8 μmol metha-
choline.
‡ FEV₁ and FEV₁/FVC were expressed as change from 100% in
ratio of observed to expected value and are adjusted for sex,
race, the natural logarithm (ln) of height, ln of age, and ln of
body mass index.
Table 3  Unadjusted and adjusted odds ratios (with 95% confidence intervals) for the association of salt intake with asthma*

<table>
<thead>
<tr>
<th>Salt score quartiles*</th>
<th>n</th>
<th>Unadjusted odds ratio</th>
<th>Adjusted odds ratio†</th>
</tr>
</thead>
<tbody>
<tr>
<td>I</td>
<td>76</td>
<td>Reference category</td>
<td>Reference category</td>
</tr>
<tr>
<td>II</td>
<td>79</td>
<td>0.90 (0.48 to 1.70)</td>
<td>1.07 (0.51 to 2.27)</td>
</tr>
<tr>
<td>III</td>
<td>82</td>
<td>0.90 (0.48 to 1.69)</td>
<td>0.81 (0.37 to 1.78)</td>
</tr>
<tr>
<td>IV</td>
<td>95</td>
<td>0.89 (0.49 to 1.62)</td>
<td>1.12 (0.54 to 2.33)</td>
</tr>
</tbody>
</table>

*Salt intake increases from quartile I to IV (see text for explanation).
†Asthma is defined as a >10% decline in FEV₁ after exercise and/or a history of diagnosed asthma.
‡Adjusted for age, sex, race, asthma in a parent, socioeconomic status, and prenatal and postnatal exposure to tobacco smoke.

Table 4  Relationship of dietary salt intake to methacholine dose-response slope (DRS)

<table>
<thead>
<tr>
<th>Salt score quartile‡</th>
<th>DRS rank coefficient (95% CI)</th>
<th>Median DRS (% per μmol)</th>
</tr>
</thead>
<tbody>
<tr>
<td>I</td>
<td>Reference category</td>
<td>5.4</td>
</tr>
<tr>
<td>II</td>
<td>5.7 (-2.9 to 36.3)</td>
<td>5.9</td>
</tr>
<tr>
<td>III</td>
<td>28.4 (-2.8 to 59.7)</td>
<td>7.7</td>
</tr>
<tr>
<td>IV</td>
<td>34.2 (4.2 to 64.2)</td>
<td>8.7</td>
</tr>
</tbody>
</table>

95% CI = 95% confidence interval.
‡Salt intake increases from quartile I to IV (see text for explanation).
§Coefficient is adjusted for age, sex, parental asthma, socioeconomic status, and prenatal and postnatal exposure to tobacco smoke.

We also investigated the relationship between usual dietary salt intake and bronchial hyperresponsiveness to methacholine by combining both cases and controls into a single group. Table 4 shows the relationship of methacholine dose response to dietary salt intake after accounting for the effects of potential confounding variables. The ranks for dose-response slope increase with increasing salt intake and are significantly different between the highest and lowest quartiles. To illustrate this, median dose-response slopes are calculated within each quartile of salt intake. There were no significant differences in these relationships according to sex, allergy skin test positivity, or race (non-significant interaction terms).

Discussion

We were unable to show any association between usual salt intake and diagnosed asthma or exercise-induced bronchospasm using a case-control design. When considering cases and controls as a single group, however, we found bronchial hyperresponsiveness to methacholine to increase with usual dietary salt intake. This discrepancy could be due to the different pathophysiological mechanisms underlying exercise and methacholine airway challenges.26

The measure of salt intake used in this study has not been validated because of the difficulty in collecting multiple 24 hour urine samples in a large number of children. The dietary method most appropriate for measuring usual dietary intake of a single nutrient is a food frequency questionnaire.14 Food frequency questionnaires for other nutrients have been used widely and found to relate well to more detailed methods of dietary evaluation.14 Misclassification of children for salt intake is likely, however, and might have resulted in the attenuation of the association between salt and asthma among our study subjects.

Since children were more likely to have been included in the subsample studied if they had a history of asthma or demonstrated exercise-induced bronchoconstriction, and because such traits are related to methacholine responsiveness, a spurious correlation between methacholine responsiveness could have been created if a history of asthma or exercise-induced bronchospasm were related to salt intake. This was not the case, however. To be sure such confounding was not present, the relation between methacholine responsiveness and salt intake was re-examined after adjusting for the occurrence of exercise-induced bronchospasm or a history of asthma and the relationship was unchanged.

The fact that more children of smoking mothers did not participate in the study could have attenuated the association between usual salt intake and asthma. This is because smoking is more prevalent in mothers with low socioeconomic status and it is possible that socioeconomic status and salt intake are related.27 At the same time, exposure to environmental tobacco smoke and airways responsiveness may be related.28 Thus, non-participation of smoking mothers will result in the preferential exclusion of children with a high intake of salt and with airways hyperresponsiveness. However, socioeconomic status and usual dietary salt intake were not found to be related in our dataset.

The biological mechanisms that might explain the role of dietary salt intake in airways responsiveness remain uncertain. As summarised by Burney, the added dietary sodium load might increase the contribution of the electronegative Na⁺ pump to the resting membrane potential and potentiate the hyperpolarisation of the bronchial muscle cells. A common aetiological role of dietary sodium to both asthma and hypertension is also suggested.2 Bronchial hyperresponsiveness has been linked to increased Na⁺/K⁺ ATPase activity in animal models,29-31 as well as to increased extracellular K⁺ and smooth muscle contraction.32-34 The control of potassium homeostasis is through the adrenergic system and this control is impaired in asthmatic subjects.35 This might explain the previously reported relationship observed between urinary potassium excretion and bronchial hyperresponsiveness.35

In conclusion, we found dietary salt intake to be related to bronchial hyperresponsiveness to methacholine, but not to childhood asthma, as determined by a history of asthma and/or a decline in FEV₁ by >10% after exercise. Our results further underline the conclusions of previous studies36 that the study of risk factors for asthma in community based studies may be substantially affected by the measure of asthma chosen, and that measures of non-specific bronchial responsiveness may be less helpful than initially hoped.
Appendix

FOOD FREQUENCY QUESTIONNAIRE

Below is a list of foods your child may eat. Please indicate how often your child eats each of the foods.

<table>
<thead>
<tr>
<th>Food</th>
<th>Most days 3-5 times per week</th>
<th>1-2 times per week</th>
<th>Less than once per week</th>
</tr>
</thead>
<tbody>
<tr>
<td>Cheese slices</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Bacon or salami or bologna</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Potato chips, peanuts, pretzels, corn chips, etc</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Salted crackers</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Canned or packaged soup</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Kraft dinner, lasagna</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Frozen TV dinners, pot pies, etc</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Dill pickles</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Ketchup, soy sauce</td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Do you like salty foods?

- Very much
- Somewhat
- Not at all

Do you usually add salt to foods?

- Yes
- No

The authors acknowledge the advice of Professors Margaret Becklake and Ben Armstrong.

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