Estimation and repeatability of the response to inhaled histamine in a community survey

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ABSTRACT Epidemiological problems arising from the absence of an agreed definition of asthma have led to the use of bronchial reactivity tests in community surveys of asthma prevalence. Since only a minority of the general population will develop bronchoconstriction in response to the dose of histamine considered acceptable for use in the community it is important to make maximum use of the data available. Several methods for summarising the information in the dose-response curve obtained from a histamine challenge test have been compared. A standardised histamine challenge test was administered to 797 subjects selected from two communities, and a repeat test to 106 subjects. The test was well accepted. For most subjects FEV<sub>1</sub> rose initially after administration of histamine (median rise 100 ml), so maximum FEV<sub>1</sub> was used as the baseline from which the 20% fall to achieve a PD<sub>20</sub> was calculated. In order to use all the data rather than just two points on the FEV<sub>1</sub>-log dose graph, PD<sub>20</sub> was estimated by means of curve fitting, and the values were compared with PD<sub>20</sub> from linear interpolation. An exponential curve was found to fit the data well. Extrapolation from the maximum dose of 4 μmol up to 8 μmol was allowed in the estimation of PD<sub>20</sub> by both methods. The curve fitting method gave slightly more reproducible PD<sub>20</sub> values than did linear interpolation, and also gave more estimates in the range 0.03–8 μmol. The repeatability of PD<sub>20</sub> compared well with that of asthmatic subjects tested in a clinical environment. Curve fitting has an advantage over linear interpolation in large community studies, for which analysis of data by computer is essential.

Introduction

Most patients with asthma show bronchial hyper-reactivity to a variety of non-allergenic stimuli, although the precise relationship between reactivity and asthma is still debated. In general, hyperreactivity is related to the severity of asthma, as judged by symptoms, requirement for treatment, and diurnal variation in peak expiratory flow. Various guidelines have recently been suggested for the standardisation of bronchial challenge tests.

The epidemiological problems arising from the absence of any agreed definition of asthma have led to interest in the possibility of using tests of bronchial reactivity in community surveys. In this situation the test needs to be quick and simple and be acceptable to the general population, and to give reproducible results under field conditions. It is essential to have a standardised protocol to be used for all subjects. Yan et al have developed a method for this purpose, giving increasing concentrations of histamine from hand held deVilbiss nebulisers and measuring the response as the dose of histamine estimated by interpolation to cause a 20% fall (PD<sub>20</sub>) in one second forced expiratory volume (FEV<sub>1</sub>). In a pilot study Britton et al showed Yan's method to be quicker and at least as reproducible as two alternative tests, and this was therefore the method chosen for the community survey.

Measurement of bronchial reactivity in epidemiological studies is subject to certain restraints. Few of the population will have had any previous experience of forced expiratory manoeuvres. Since the general population may be less tolerant of side effects than subjects in the laboratory, the maximum dose of histamine that can be given is lower, and a smaller proportion of subjects will have a PD<sub>20</sub> that can be esti-
mated. Hence it is important to obtain the maximum information from the data that are obtained.

In the main study Yan's method of histamine challenge was used to measure PD_{20} in 797 subjects as part of a community survey of asthma prevalence in two areas of southern England. Repeat studies were performed in 106 subjects 1–14 days after the first test. In the study reported in this paper we set out to establish the best way to analyse histamine challenge dose-response data obtained in an epidemiological setting by determining: (1) the proportion of subjects for whom a PD_{20} value was obtained by fitting dose-response curves compared with the use of linear interpolation; (2) the repeatability of PD_{20} values for the two methods; (3) whether PD_{20} should be estimated as a 20% fall from the post-saline FEV_{1} or the maximum FEV_{1}.

PD_{10} was estimated and assessed in the same way as PD_{20}. The fact that more subjects would have a PD_{10} than a PD_{20} value of ≤8 \mu mol could favour the use of PD_{10} in a community survey if the results are as repeatable.

**Methods**

**Subjects**

A questionnaire about symptoms of asthma was administered to all subjects aged 18–64 years in two villages in Hampshire and a market town in Dorset as part of a study of asthma prevalence. After completion of the questionnaire 1325 subjects were asked to attend for a histamine challenge test at their general practitioner's surgery or health clinic. The subjects consisted of two groups, a 20% random sample (855) of the 4277 subjects aged 18–64 years who returned a completed questionnaire and all remaining subjects (470) who answered "Yes" to the question "Have you had wheezing or whistling in your chest at any time in the last 12 months?" In this paper no distinction is made between the two groups of subjects; the prevalence of hyperreactivity and its relation to smoking history and skin sensitivity as estimated from the random sample has been reported and the relation of hyperreactivity to associated symptoms will be reported later.

Subjects who had taken theophyllines or anti-histamines in the previous 24 hours or a bronchodilator in the last six hours were asked to return later after omitting this treatment. Ethical approval was obtained from the local ethical committees and all subjects had the test explained and signed a consent form before the test was carried out.

A total of 170 subjects were invited to return for a second histamine challenge test. These consisted of a 10% random sample of all subjects in one area and in addition all subjects whose FEV_{1} had fallen by 20% in the initial test in both areas, plus a few subjects whose fall in FEV_{1} approached 20%. The decision to administer the second test by the same or a different person was taken at random.

**Measurements**

Height was measured and predicted FEV_{1} calculated as recommended by Cotes et al. \(^9\) FEV_{1} and forced vital capacity were measured with a dry spirometer (Vitalograph, Buckingham, England). Initial FEV_{1} was recorded as the maximum of three consecutive measurements that agreed to within 5%. Subjects whose initial FEV_{1} was less than 60% of the predicted value were not challenged with histamine.

Other details of the histamine challenge test were as described by Yan et al. \(^6\) with doubling doses of from 0.03 to 4 \mu mol administered to subjects with a history of wheezing or whose post-saline FEV_{1} was less than 90% predicted. All other subjects were given 0.06 \mu mol histamine followed by quadrupling doses of histamine until their FEV_{1} had fallen by at least 10%, when the challenge regimen was changed to the slower schedule. The test was stopped when the FEV_{1} had fallen by 20% or more from the post-saline value or the 40 \mu mol dose had been given, or at the subject's request.

**Estimation of PD_{20}**

PD_{20} FEV_{1} was estimated by linear interpolation on the basis of data from the last two doses administered,\(^10\) with extrapolation to one doubling dose (8 \mu mol) beyond the maximum administered, following Cockcroft et al. \(^11\) PD_{20} was estimated by using all the available data and fitting the exponential curve

\[
\log (c - y) = a + bx, \quad \text{(Curve 1)}
\]

where \(y\) is FEV_{1}, \(x\) is \log_{10} (dose), \(c\) represents mean FEV_{1} before it is affected by histamine, \(b\) is a "slope" parameter, and \(a\) is a curve position parameter. Since Woolcock et al.\(^12\) fitted a logistic curve, we also fitted

\[
\log (c - y)/y = a + bx. \quad \text{(Curve 2)}
\]

Extrapolation to a dose of 8 \mu mol was allowed for both curves as with the linear interpolation method. Curves were fitted provided that two or more doses of histamine were administered—that is, they were fitted to all sets of data for which PD_{20} values were obtained by linear interpolation. When only two doses of histamine were given the post-saline FEV_{1} was used as the estimate of \(c\). Details are given in the appendix.

Most subjects showed a small increase in FEV_{1} after the low dose of histamine, so we also calculated the estimates described above as the dose producing a 20% fall from maximum FEV_{1}. Linear interpolation uses the doses either side of the estimate. Occasionally both the last two doses administered produced a
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The methods of estimating PD$_{20}$ were compared according to the number of subjects who achieved a PD$_{20}$ and the repeatability of the measurements.

PD$_{10}$ was estimated, by linear interpolation and from the same curves, in a similar way to PD$_{20}$.

**Repeatability**

The mean and standard deviation of the difference between the estimates of PD$_{20}$ for repeat tests were calculated for the subjects tested twice. Repeatability was also calculated for PD$_{10}$, post-saline FEV$_1$, and maximum FEV$_1$. The within subject standard deviation was calculated by dividing the standard deviation of the difference of two repeat estimates by the square root of 2.$^{13}$ A dimensionless measure of repeatability, intraclass correlation coefficient,$^{14}$ was also calculated.

**Results**

Of 1325 subjects invited to undergo a histamine challenge test, 834 (63%) agreed. Thirty seven subjects were not given histamine, four owing to difficulty in complying with instructions and 33 because the initial or post-saline FEV$_1$ was less than 60% predicted. Thus 797 subjects were given at least one dose of histamine at the first test, including four subjects with an initial FEV$_1$ less than 60% predicted who were given histamine in error; of these, 512 (64.2%) were randomly selected.

Most subjects (74.4%) increased their FEV$_1$ after the post-saline measurement, 83 (10.4%) having their maximum recorded FEV$_1$ after the final 4 μmol dose of histamine. The increase in FEV$_1$ over the post-saline FEV$_1$ ranged up to 0.9 litre in absolute terms (a 30% increase), although the median increase was only 100 ml.

Four subjects were given just one dose of histamine, of whom three had a 20% fall in FEV$_1$ after that dose and so did not have a PD$_{20}$ that could be estimated. Each method of estimation of PD$_{20}$ was applied to the data for the remaining 793 subjects. The residual standard deviation about the curves was less than 0.21 for 88.1% of subjects for curve 1 and 86.0% for curve 2; the mean residual standard deviation was 0.151 for each curve. The larger residual standard deviations were found to be due to random variation in the FEV$_1$ rather than systematic deviations from the fitted curves. This is illustrated in the example in figure 1, where the residual standard deviations about the curves were 0.23 and 0.211.
NUMBERS OF SUBJECTS WITH PD_{10} AND PD_{20} VALUES

The number of estimates of PD_{20} is shown in table 1, and varies according to method from 20-2% to 26-2% of the 793 subjects. Curve fitting produced more estimates in the range 0-03–8 μmol than did linear interpolation, the difference being between 28 and 38 subjects. Use of the maximum FEV_{1} as baseline rather than the post-saline FEV_{1} increased the number of estimates, since a 20% fall was achieved at greater absolute FEV_{1}. The numbers achieving a 20% fall from the maximum FEV_{1}, when calculated by linear interpolation and by fitting curves 1 and 2, were 173, 208, and 201 respectively, whereas the corresponding numbers achieving a 10% fall were 292, 324, and 319, about 50% more.

REPEATABILITY

A total of 112 (66%) of the 170 subjects invited presented for retest. These included 19 (48%) of the 40 randomly selected subjects and 93 (72%) of the 130 who had a fall of 20%, or nearly 20%, in FEV_{1} at the first test. Six subjects were not given histamine on the second occasion, five because of recurrent wheezing after the first test, and one because of error in calculating the initial FEV_{1} as being less than 60% predicted. Two subjects had a 20% fall in FEV_{1} after one dose of histamine, so that a second PD_{20} could not be estimated. Duplicate estimates of PD_{20} were therefore available in 104 subjects.

The repeatability of log PD_{20} values obtained by each method is given in table 2 as the standard deviation of the differences between test and retest values. This is given for all 104 subjects, for the 90 subjects with at least one estimate less than or equal to 8 μmol, and for the 64 subjects with all estimates less than 8 μmol. PD_{20} values estimated by curve fitting were a little more repeatable (had a smaller standard deviation) than those for linear interpolation, and PD_{20} values from curve 1 were slightly more repeatable than those from curve 2 (columns one and two). The repeatability of the three methods was similar when restricted to the 64 subjects with no estimate greater than 8 μmol. The number with a PD_{20} of 8 μmol or less on both occasions varied from 68 for linear interpolation, with post-saline FEV_{1} as baseline, to 73 for curve 1. Figure 2 shows the relation between curve 1 PD_{20} for the second test to that for the first, on the basis of fall from maximum FEV_{1}. The choice of baseline had little influence on the repeatability of the estimated PD_{20}.

The post-saline FEV_{1} was slightly less repeatable than maximum FEV_{1}. Table 3 also shows the repeatability of PD_{10} estimates from curve 1, with FEV_{1}

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**Table 1** Numbers of censored and estimated PD_{20} values obtained by methods of linear interpolation and curve fitting for the 793 subjects given at least two doses of histamine

<table>
<thead>
<tr>
<th>Method</th>
<th>Censored &lt;0.03 μmol</th>
<th>Estimated</th>
<th>Censored &gt;8 μmol</th>
</tr>
</thead>
<tbody>
<tr>
<td>(a) PD_{20} is dose producing 20% fall in FEV_{1} from post-saline FEV_{1}</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Linear interpolation</td>
<td>1</td>
<td>160</td>
<td>632</td>
</tr>
<tr>
<td>Curve 2</td>
<td>2</td>
<td>198</td>
<td>593</td>
</tr>
<tr>
<td>Curve 2</td>
<td>2</td>
<td>190</td>
<td>601</td>
</tr>
<tr>
<td>(b) PD_{20} is dose producing 20% fall in FEV_{1} from maximum FEV_{1}</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Linear interpolation</td>
<td>1</td>
<td>173</td>
<td>619</td>
</tr>
<tr>
<td>Curve 1</td>
<td>2</td>
<td>208</td>
<td>583</td>
</tr>
<tr>
<td>Curve 2</td>
<td>2</td>
<td>201</td>
<td>590</td>
</tr>
</tbody>
</table>

PD_{20}—histamine dose (μmol) causing 20% fall in FEV_{1}.

**Table 2** Repeatability of log_{10} PD_{20} (values greater than log 8 set to log 8 where appropriate)

<table>
<thead>
<tr>
<th>Difference between 2nd and 1st test (log-scale)</th>
<th>All subjects tested twice (n = 104)</th>
<th>Subjects with at least one value ≤8 μmol (n = 90)</th>
<th>Subjects with all values ≤8 μmol (n = 64)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Mean</td>
<td>SD</td>
<td>Mean</td>
<td>SD</td>
</tr>
<tr>
<td>(a) PD_{20} is dose producing 20% fall in FEV_{1} from post-saline FEV_{1}</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Linear interpolation</td>
<td>-0.04</td>
<td>0.40</td>
<td>-0.04</td>
</tr>
<tr>
<td>Curve 1</td>
<td>-0.03</td>
<td>0.36</td>
<td>-0.03</td>
</tr>
<tr>
<td>Curve 2</td>
<td>-0.02</td>
<td>0.39</td>
<td>-0.02</td>
</tr>
<tr>
<td>(b) PD_{20} is dose producing 20% fall in FEV_{1} from maximum FEV_{1}</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Linear interpolation</td>
<td>-0.03</td>
<td>0.39</td>
<td>-0.03</td>
</tr>
<tr>
<td>Curve 1</td>
<td>-0.02</td>
<td>0.36</td>
<td>-0.03</td>
</tr>
<tr>
<td>Curve 2</td>
<td>-0.01</td>
<td>0.39</td>
<td>-0.01</td>
</tr>
</tbody>
</table>

PD_{20}—histamine dose (μmol) causing 20% fall in FEV_{1}.
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Table 3  Repeatability of FEV₁ and curve parameters

<table>
<thead>
<tr>
<th></th>
<th>2nd and 1st test</th>
<th>Within subject</th>
<th>Between subject</th>
<th>Intra-class correlation coefficient</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Mean</td>
<td>SD</td>
<td>SD</td>
<td>SD</td>
</tr>
<tr>
<td>FEV₁ post-saline (litres)</td>
<td>107</td>
<td>0·01</td>
<td>0·42</td>
<td>0·30</td>
</tr>
<tr>
<td>FEV₁ maximum (litres)</td>
<td>106</td>
<td>0·02</td>
<td>0·40</td>
<td>0·28</td>
</tr>
<tr>
<td>log₁₀(PD₁₀)</td>
<td>(estimated as dose producing 20% fall from FEV₁ maximum, on the basis of curve 1)</td>
<td>90</td>
<td>-0·03</td>
<td>0·39</td>
</tr>
<tr>
<td>log₁₀(PD₁₀)</td>
<td>(estimated as dose producing 10% fall from FEV₁ maximum, on the basis of curve 1)</td>
<td>90</td>
<td>0·01</td>
<td>0·46</td>
</tr>
</tbody>
</table>

Fig 2  Relation between (log) PD₂₀ (histamine dose (µmol) causing 20% fall in FEV₁) at the second test to (log) PD₂₀ at the first test, as estimated from curve 1 with maximum FEV₁ as baseline. The line — — — — is the line of identity and the lines — — — denote a difference of one doubling dose of histamine.

maximum as baseline, compared with that of PD₂₀ in the same 90 subjects, whose data are given in column 2 of table 2. The intra-class correlation coefficient was less than that for PD₂₀ in these subjects. Estimates of repeatability of PD₁₀ from curve 2, and estimates based on post-saline FEV₁, were similar. PD₁₀ estimated by linear interpolation was less repeatable, the difference in repeatability being similar to that of PD₂₀ (data available but not shown).

Repeatability was affected to a small extent by differences in time of day of measurement, the time between tests (range 1–14 days), and the person performing the test. Data were available for the 27 subjects retested one day later, at the same time of day to within one hour, and by the same tester. If we excluded two subjects with a PD₂₀ greater than 8 µmol on both occasions, the within subject standard deviation of curve 1 PD₂₀ was reduced from 0·27 (table 3) to 0·18.

Discussion

Studies on asthma prevalence have been greatly hampered by the absence of an agreed definition of asthma. Measurement of bronchial responsiveness provides an objective test, and since it is an important marker of asthma there has been interest recently in exploring the role of bronchial reactivity measurement in the assessment of asthma prevalence in the community.

Measurement of bronchial reactivity in large surveys poses different problems from measurements in the laboratory. The incidence of side effects considered acceptable in a community may be lower than for subjects in the laboratory, so a lower maximum dose of histamine has to be given. We found in a pilot study that the test described by Yan et al was feasible in inexperienced subjects and that 4 µmol histamine was the highest dose that was generally well tolerated. With this low maximum dose of histamine fewer subjects will develop bronchoconstriction, so it is important to estimate the response by a method that makes maximum use of the information available.

PD₂₀ values have usually been obtained by linear interpolation between the response to the last two doses of histamine, with extrapolation up to one further doubling concentration of histamine by some workers. This technique is simple and gives an immediate answer. Fitting a curve to the data is more complicated but, when the data are of necessity analysed by computer, curve fitting is no more difficult than linear interpolation. Since we needed to analyse a large number of histamine dose-response curves we decided to compare the linear interpolation method with two curve fitting models, extrapolating up to one doubling dose (8 µmol) for each method to increase the number of estimates. We did not include in this comparison estimates of PD₂₀ if only one dose of histamine was given, as the published method of estimation requires linear interpolation between post-saline FEV₁ and post-histamine FEV₁ at the first dose on a linear scale and is inappropriate when linear
Interpolation is otherwise carried out on a log-dose scale (post-saline “dose” \( = 0 \), and \( \log (0) = -\infty \)).

We looked at two curves, both of which are described by the minimum three parameters, upper asymptote, slope, and position. Both curves appear to be preferable to the quadratic equation recommended by Neijen’s \(^{15}\) which is less flexible in shape and has no slope parameter. Curve 1 was the simplest such curve. Curve 2 was chosen because Woolcock et al.\(^ {12}\) had fitted a logistic curve to their data. The repeatability of PD\(_{20}\) values from curve 2 was less good than that of values from curve 1.

On the basis of fall from maximum FEV\(_1\), a PD\(_{20}\) value was obtained in 173 subjects by the linear interpolation method. Curve fitting increased the number of subjects with an estimated PD\(_{20}\) value by 35 and 28 subjects respectively for curves 1 and 2. The lower number from linear interpolation occurred in part because an extrapolated value was obtained only if the FEV\(_1\) after the final dose was less than the previous FEV\(_1\), whereas curve fitting provided an estimate whether or not this was the case. In the subjects tested twice there were correspondingly more estimates by curve fitting than by linear interpolation, which led to more subjects with one or both estimates less than 8 µmol. The repeatability of PD\(_{20}\) values obtained by fitting curve 1 was better than that of values obtained by linear interpolation, as shown by column 2 of table 2. Column 3, which contains results only for subjects with all estimates less than 8 µmol, apparently shows greater repeatability for linear interpolation and curve 2 estimates than does column 2 and lower repeatability for curve 1 estimates. This is because some of the subjects who are included in column 2 but not in column 3 had much more discrepant values derived from linear interpolation and curve 2 than from curve 1, and column 3 is included only to demonstrate this point. Selective exclusion of subjects with values above 8 µmol for each method would be even more misleading as different subjects would be included for each method.

With any measurement having an upper limit above which values cannot be obtained there is no ideal solution to the problem of estimating repeatability. Methods can be properly compared only when carried out on the same or comparable subjects, and even then care must be taken not to exclude subjects more variable by one method than another. There is also the problem of whether the between subject variation used in the calculation of the intraclass correlation coefficient should be estimated from subjects tested twice or from a sample more representative of the population. PD\(_{10}\) had a greater within subject variation and greater between subject variation than PD\(_{20}\) in the 90 subjects; the intraclass correlation coefficient showed PD\(_{10}\) to be relatively less repeatable than PD\(_{20}\). Judged against variation in all 793 subjects studied, PD\(_{10}\) and PD\(_{20}\) had similar repeatability. We have preferred to use PD\(_{20}\) since this has been used more often by others.

Other measures of reactivity that have been used elsewhere include threshold dose,\(^ {16}\) which has been shown to be less repeatable\(^ {17}\) than PD\(_{20}\) and to add little information.

Most subjects (75%) increased their FEV\(_1\) after the post-saline measurement. In most the increase was less than 100 ml, which is within the 200 ml 95% range that would be expected in subjects unfamiliar with a spirometer.\(^ {17}\) If we assume that the highest value is closer to the true FEV\(_1\) it is appropriate to take the maximum FEV\(_1\) as a baseline rather than the post-saline FEV\(_1\). This decision is supported by the slightly smaller intraclass correlation coefficient of the maximum FEV\(_1\) than of the post-saline FEV\(_1\) (table 3). The difference, however, between the results from the two methods is trivial, the PD\(_{20}\) derived from the maximum FEV\(_1\) being on average only 0·008 log µmol (equivalent to 0·03 doubling doses) less than that derived from the post-saline FEV\(_1\) (standard deviation of differences 0·036 log µmol).

The repeatability of PD\(_{20}\) in this study, as estimated from curve 1, was similar to that found in laboratory based studies. Dehaut et al.\(^ {18}\) reported a 95% single determination “confidence interval” (strictly speaking a range, not a confidence interval) of 1·59 doubling concentrations; our within subject standard deviation of 0·27 gives a corresponding figure of 1·79 doubling doses. When our calculation was restricted to data tested under conditions more like those that would be imposed in a laboratory setting the within subject standard deviation was 0·18, equivalent to a 95% single determination interval of 1·2 doubling doses. Other laboratory studies have found better agreement than Dehaut between duplicate measurements, in trained and selected subjects; but these results should not be compared with our findings. An epidemiological study of a community population must include all respondents irrespective of technique if bias is to be avoided, and this will inevitably decrease repeatability of the method used.

Variability in FEV\(_1\) makes linear interpolation from just two measurements unreliable and also reduces the number of estimates; the FEV\(_1\) at 4 µmol can, by chance, be greater than the previous value; this renders linear extrapolation to a PD\(_{20}\) impossible whereas curve fitting will give an estimate. We did not obtain a direct estimate of the short term repeatability of FEV\(_1\), but the residual standard deviation about curve 1 was 0·151, well within the range of 0·1–0·31 for different subgroups quoted by Tweeddale et al.\(^ {18}\) for repeat FEV\(_1\) measurements in inexperienced subjects. Extrapolation using curve 1 increased the number of
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PD$_{20}$ values by about 20% without a reduction in repeatability (data available but not included). Extrapolation to one doubling dose beyond the maximum administered by fitting a curve to all the data is therefore justified and desirable in epidemiological surveys in which the maximum dose of histamine is low. In clinical studies, where larger doses can be given and where ease of computation may be important, interpolation between the last two points gives satisfactory results.

We would like to thank Mrs Renee Witts, Miss Marcia Kelson, Miss Karen Snowden, and Miss Leslie Neville for their assistance with the survey; Mr Frank Anderson and Mr Douglas Corfield for their help with testing lung function; Drs TE Roberts, P Harker, and D Law for helping to set up the study in the first instance; and Drs TE Wilson, D Davies, J Davies, D Pollard and R Lorge and the group practices in the areas for their cooperation with the study. We are greatly indebted to Mrs A Childs for preparing the manuscript. Finally we would like to thank the Department of Health and Social Security for financial support.

Appendix: Fitting the curves

Curve 1 is an exponential curve with a “slope” ($b$) and an upper asymptote ($c$) provided that $b$ is positive.

Curve 2 is a logistic (sigmoid) curve with an upper and a lower asymptote. The lower asymptote was fixed at zero FEV$_1$ as doses of histamine were too low for FEV$_1$ to reach the asymptote. (Woolcock et al.,$^{12}$ who gave higher doses, attempted to lower the asymptote, but fixed the upper asymptote at the post-saline FEV$_1$.)

Each curve was fitted iteratively by a two step procedure, with the post-saline FEV$_1$ as the initial estimate of parameter $c$. Step 1 estimates $a$ and $b$. Given $a$ and $b$, step 2 estimates $c$ to give the minimum residual sum of squares on the FEV$_1$ scale. Iteration continued until there was a reduction in the residual sum of squares from the previous iteration and:

(i) $c$ changed by less than 0-01;

or

(ii) $c$ changed by less than 0-05 and number of iterations $> 10$;

or

(iii) number of iterations $> 50$.

**Curve 1**

$y = c - \exp (a + bx)$

**Step 1** Estimate $a$ and $b$ by regression of log$_{10}$ (c - y) on x. If some values of FEV$_1$ (y) are greater than or equal to c, only data for values of x greater than the highest x for which $y > c$ are used.

**Step 2** Estimate $c = (\Sigma y + \Sigma \exp (a + bx))/n$, where n is the total number of doses of histamine.

**Curve 2**

$y = c/(1 + \exp (a + bx))$

**Step 1** Estimate $a$ and $b$ by regression of log$_{10}$ (c - y)/y on x. Omit data as for curve 1.

**Step 2** Estimate $c = \Sigma yz/\Sigma^2$, where $z = 1/(1 + \exp (a + bx))/n$.

Because minimisation in step 1 is on the log scale and that in step 2 on the original scale, non-convergence can occur. Omission of values in step 1 can also cause this, and the second and third criteria for stopping iteration were to cope with the few data sets for which this occurred. As rising curves were of no interest, $b$ was set to zero if a value less than zero resulted from step 1.

Parameter $c$ is the asymptote of the curve unless $b$ is zero, when the “curve” is a straight line with ordinate $c - \exp b$ for curve 1 or $c/(1 + \exp b)$ for curve 2. The algorithms are approximate in that in some cases the minimum residual sum of squares will not be found—generally when the curve fits the data poorly. The computing time required for algorithms guaranteeing the best fit for each curve to each of the data sets, which totalled 897, could not be justified. The algorithms failed to converge for less than 5% of data sets.

The program to fit the curves was written in FORTRAN for a minicomputer. The calculations for fitting curve 1 could be carried out on a microcomputer, or even on a sophisticated programmable calculator provided that sufficient data stores as well as program steps are available.

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